

1. Rosenbrock 函数 (香蕉函数)

我们要采样的后验分布满足

$$\begin{aligned}\rho_{\text{post}}(\theta) &= e^{-\Phi_R(\theta)} \\ \Phi_R &= \frac{1}{2} \left((y - \mathcal{G}(\theta))^T \Sigma_{\eta}^{-1} (y - \mathcal{G}(\theta)) + (\theta - r_0)^T \Sigma_0^{-1} (\theta - r_0) \right) \\ &= \frac{1}{2} \left(\frac{100(\theta_2 - c_1 \theta_1^2)^2}{c_2} + \frac{(1 - \theta_1)^2}{c_2} + \frac{\theta_1^2}{100} + \frac{\theta_2^2}{100} \right)\end{aligned}$$

暴力网格搜索

In [3]:

```
using PyPlot

function Phi_R_Rosenbrock(theta1, theta2, c1, c2)
    return (100*(theta2 - c1*theta1^2)^2/c2 + (1.0 - theta1)^2/c2 + theta1^2/100 + theta2^2/100)/2.0
end

Lx, Ux = -4.0, 8.0
Ly, Uy = -2.0, 10#3.5
N = 1000
X = zeros(N, N)
Y = zeros(N, N)
rho = zeros(N, N)
for i = 1:N
    for j = 1:N
        X[i, j], Y[i, j] = Lx + (Ux - Lx) * (i-1)/(N-1), Ly + (Uy - Ly) * (j-1)/(N-1)
    end
end
dx = dy = 1/(N-1)

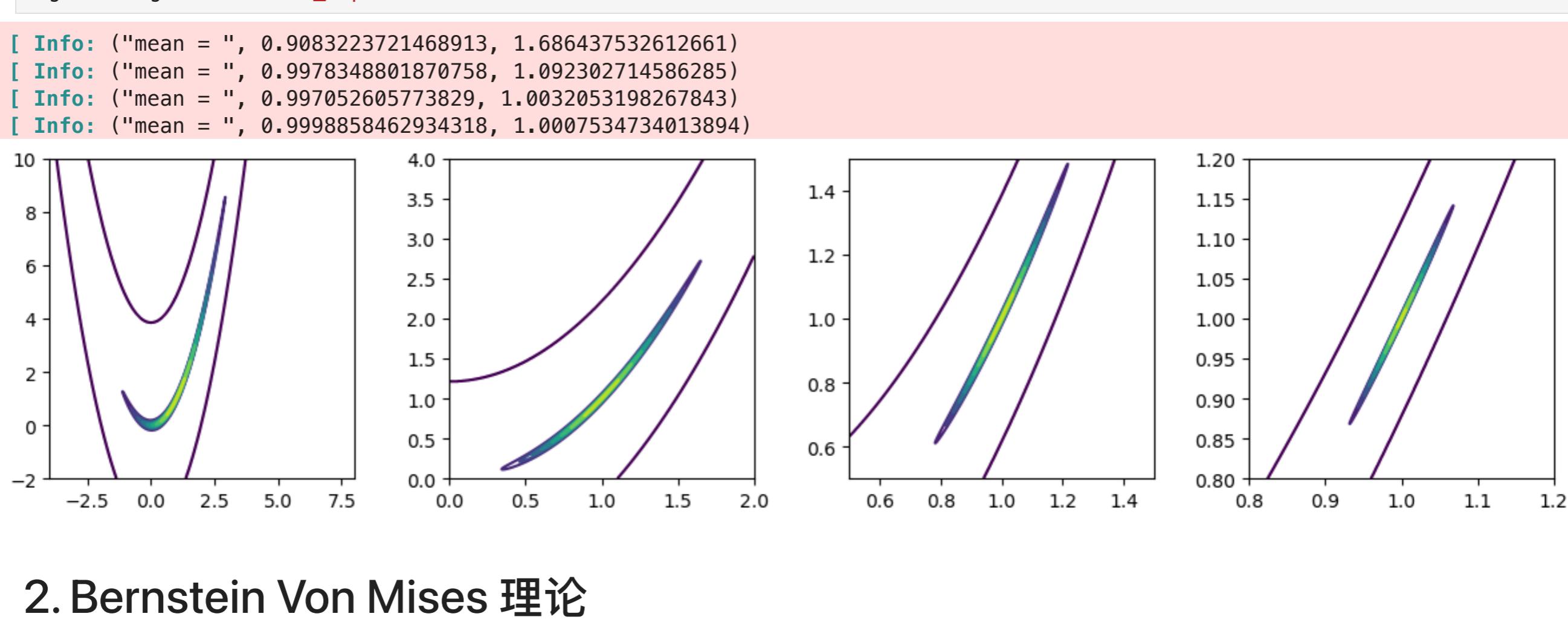
fig, ax = PyPlot.subplots(ncols=4, nrows=1, sharex=false, sharey=false, figsize=(12,3))

for k = 1:4
    c1, c2=10^(-(k-1.0)), 1.0
    for i = 1:N
        for j = 1:N
            rho[i, j] = Phi_R_Rosenbrock(X[i, j], Y[i, j], c1, c2)
        end
    end

    rho .= exp.(-rho)
    Z = sum(rho)
    rho ./= Z
    rho ./= (dx * dy)
    ax[k].contour(X, Y, rho, 10)
    @info "mean = ", sum(X.*rho)/sum(rho), sum(Y.*rho)/sum(rho)
end

fig.tight_layout()
fig.savefig("Rosenbrock_1.pdf")
```

```
[ Info: ("mean = ", 0.9083223721468913, 1.686437532612661)
[ Info: ("mean = ", 0.9893276577929085, 0.19675099271002844)
[ Info: ("mean = ", 0.9900925963862894, 0.019701671046172844)
[ Info: ("mean = ", 0.9901002636870367, 0.001970193755655862)
```



In [4]:

```
Lx, Ux = [-4.0;0.0;0.5;0.8], [8.0;2.0;1.5;1.2]
Ly, Uy = [-2.0;0.0;0.5;0.8], [10.0;4.0;1.5;1.2]
N = 1000
X = zeros(N, N)
Y = zeros(N, N)
rho = zeros(N, N)
for i = 1:N
    for j = 1:N
        end
    end

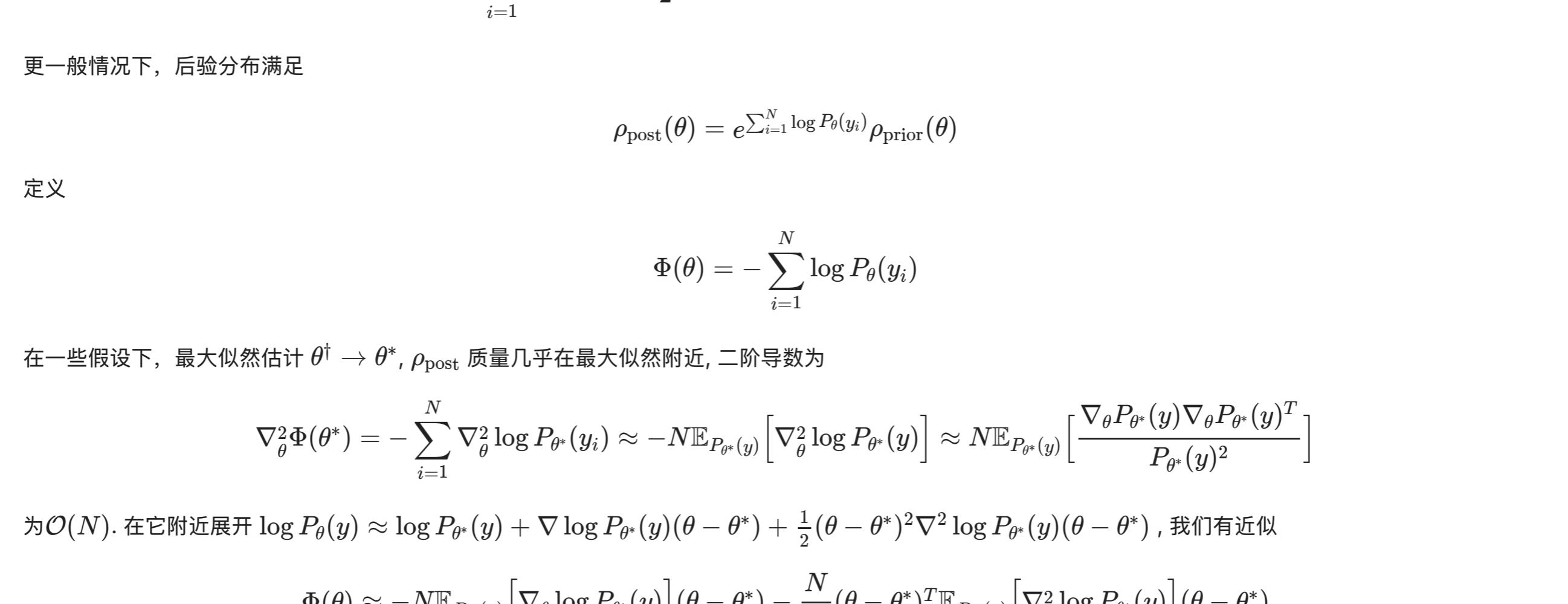
fig, ax = PyPlot.subplots(ncols=4, nrows=1, sharex=false, sharey=false, figsize=(12,3))

for k = 1:4
    c1, c2= 1.0, 10^(-(k-1.0))
    for i = 1:N
        for j = 1:N
            X[i, j], Y[i, j] = Lx[k] + (Ux[k] - Lx[k]) * (i-1)/(N-1), Ly[k] + (Uy[k] - Ly[k]) * (j-1)/(N-1)
            rho[i, j] = Phi_R_Rosenbrock(X[i, j], Y[i, j], c1, c2)
        end
    end

    rho .= exp.(-rho)
    Z = sum(rho)
    rho ./= Z
    rho ./= (dx * dy)
    ax[k].contour(X, Y, rho, 10)
    @info "mean = ", sum(X.*rho)/sum(rho), sum(Y.*rho)/sum(rho)
end

fig.tight_layout()
fig.savefig("Rosenbrock_2.pdf")
```

```
[ Info: ("mean = ", 0.9083223721468913, 1.686437532612661)
[ Info: ("mean = ", 0.9978348801870758, 1.092302714586285)
[ Info: ("mean = ", 0.997052605773829, 1.0032053198267843)
[ Info: ("mean = ", 0.9998858462934318, 1.0007534734013894)
```



2. Bernstein Von Mises 理论

小噪音情况

对于后验分布，我们有

$$\rho_{\text{post}}(\theta) = e^{-\Phi(\theta)} \rho_{\text{prior}}(\theta) \quad \Phi(\theta) = \frac{1}{2\gamma^2} (y - \mathcal{G}(\theta))^T \Sigma_{\eta}^{-1} (y - \mathcal{G}(\theta))$$

假设 $\Phi(\theta)$ 的极小值是 θ^* ，那么当 $\gamma \rightarrow 0$

$$\frac{\rho_{\text{post}}(\theta^*)}{\rho_{\text{post}}(\theta)} = e^{\Phi(\theta^*) - \Phi(\theta)} \frac{\rho_{\text{prior}}(\theta^*)}{\rho_{\text{prior}}(\theta)} \rightarrow \infty$$

质量将集中在 θ^* 附近 $|\theta - \theta^*| < \gamma$ ，我们可以线性化，

$$\Phi(\theta) \approx \frac{1}{2\gamma^2} (y - \mathcal{G}(\theta^*) - \nabla \mathcal{G}(\theta^*)^T (\theta - \theta^*)) \Sigma_{\eta}^{-1} (y - \mathcal{G}(\theta^*) - \nabla \mathcal{G}(\theta^*) (\theta - \theta^*)) = \Phi(\theta^*) + \frac{1}{2\gamma^2} (\theta - \theta^*)^T \nabla \mathcal{G}(\theta^*)^T \Sigma_{\eta}^{-1} \nabla \mathcal{G}(\theta^*) (\theta - \theta^*)$$

转化为线性贝叶斯问题处理。

大量数据情况

数据足够多，也可以转化成误差足够小。对于数据足够多的情况，后验分布满足

$$\rho_{\text{post}}(\theta) = e^{-\sum_{i=1}^N \Phi(\theta; y_i)} \rho_{\text{prior}}(\theta) \quad \Phi(\theta; y) = \frac{1}{2} (y - \mathcal{G}(\theta))^T \Sigma_{\eta}^{-1} (y - \mathcal{G}(\theta))$$

假设 $y_i = \mathcal{G}(\theta^*) + \eta_i$

$$\sum_{i=1}^N \Phi(\theta; y_i) = \frac{1}{2} \sum_{i=1}^N (\mathcal{G}(\theta^*) + \eta_i - \mathcal{G}(\theta))^T \Sigma_{\eta}^{-1} (\mathcal{G}(\theta^*) + \eta_i - \mathcal{G}(\theta))$$

在一些假设下，最大似然估计 $\theta^* \rightarrow \theta^*$, ρ_{post} 质量几乎在最大似然附近，二阶导数为 $\mathcal{O}(N)$. 在它附近展开 $\mathcal{G}(\theta) \approx \mathcal{G}(\theta^*) + \nabla \mathcal{G}(\theta^*)(\theta - \theta^*)$ ，我们有近似

$$\sum_{i=1}^N \Phi(\theta; y_i) \approx \frac{N}{2} (\theta - \theta^*)^T \nabla \mathcal{G}(\theta^*) \Sigma_{\eta}^{-1} \nabla \mathcal{G}(\theta^*) (\theta - \theta^*)$$

更一般情况下，后验分布满足

$$\rho_{\text{post}}(\theta) = e^{\sum_{i=1}^N \log P_{\theta}(y_i)} \rho_{\text{prior}}(\theta)$$

定义

$$\Phi(\theta) = -\sum_{i=1}^N \log P_{\theta}(y_i)$$

在一些假设下，最大似然估计 $\theta^* \rightarrow \theta^*$, ρ_{post} 质量几乎在最大似然附近，二阶导数为

$$\nabla_{\theta}^2 \Phi(\theta^*) = -\sum_{i=1}^N \nabla_{\theta}^2 \log P_{\theta^*}(y_i) \approx -N \mathbb{E}_{P_{\theta^*}(y)} [\nabla_{\theta}^2 \log P_{\theta^*}(y)] \approx N \mathbb{E}_{P_{\theta^*}(y)} \left[\frac{\nabla_{\theta} P_{\theta^*}(y) \nabla_{\theta} P_{\theta^*}(y)^T}{P_{\theta^*}(y)^2} \right]$$

为 $\mathcal{O}(N)$. 在它附近展开 $\log P_{\theta}(y) \approx \log P_{\theta^*}(y) + \nabla \log P_{\theta^*}(y)(\theta - \theta^*) + \frac{1}{2} (\theta - \theta^*)^2 \nabla^2 \log P_{\theta^*}(y)(\theta - \theta^*)$ ，我们有近似

$$\begin{aligned}\Phi(\theta) &\approx -N \mathbb{E}_{P_{\theta^*}(y)} [\nabla_{\theta} \log P_{\theta^*}(y)] (\theta - \theta^*) - \frac{N}{2} (\theta - \theta^*)^T \mathbb{E}_{P_{\theta^*}(y)} [\nabla_{\theta}^2 \log P_{\theta^*}(y)] (\theta - \theta^*) \\ &\approx \frac{N}{2} (\theta - \theta^*)^T I(\theta^*) (\theta - \theta^*)\end{aligned}$$

2. 多峰函数

我们要采样的后验分布满足

$$\begin{aligned}\rho_{\text{post}}(\theta) &= e^{-\Phi_R(\theta)} \\ \Phi_R &= \frac{1}{2} \left((y - \mathcal{G}(\theta))^T \Sigma_{\eta}^{-1} (y - \mathcal{G}(\theta)) + (\theta - r_0)^T \Sigma_0^{-1} (\theta - r_0) \right) \\ &= \frac{1}{2} \left((4 - (\theta_2 - \theta_1)^2)^2 + (c - \theta_1)^2 + \theta_2^2 \right)\end{aligned}$$

In [5]:

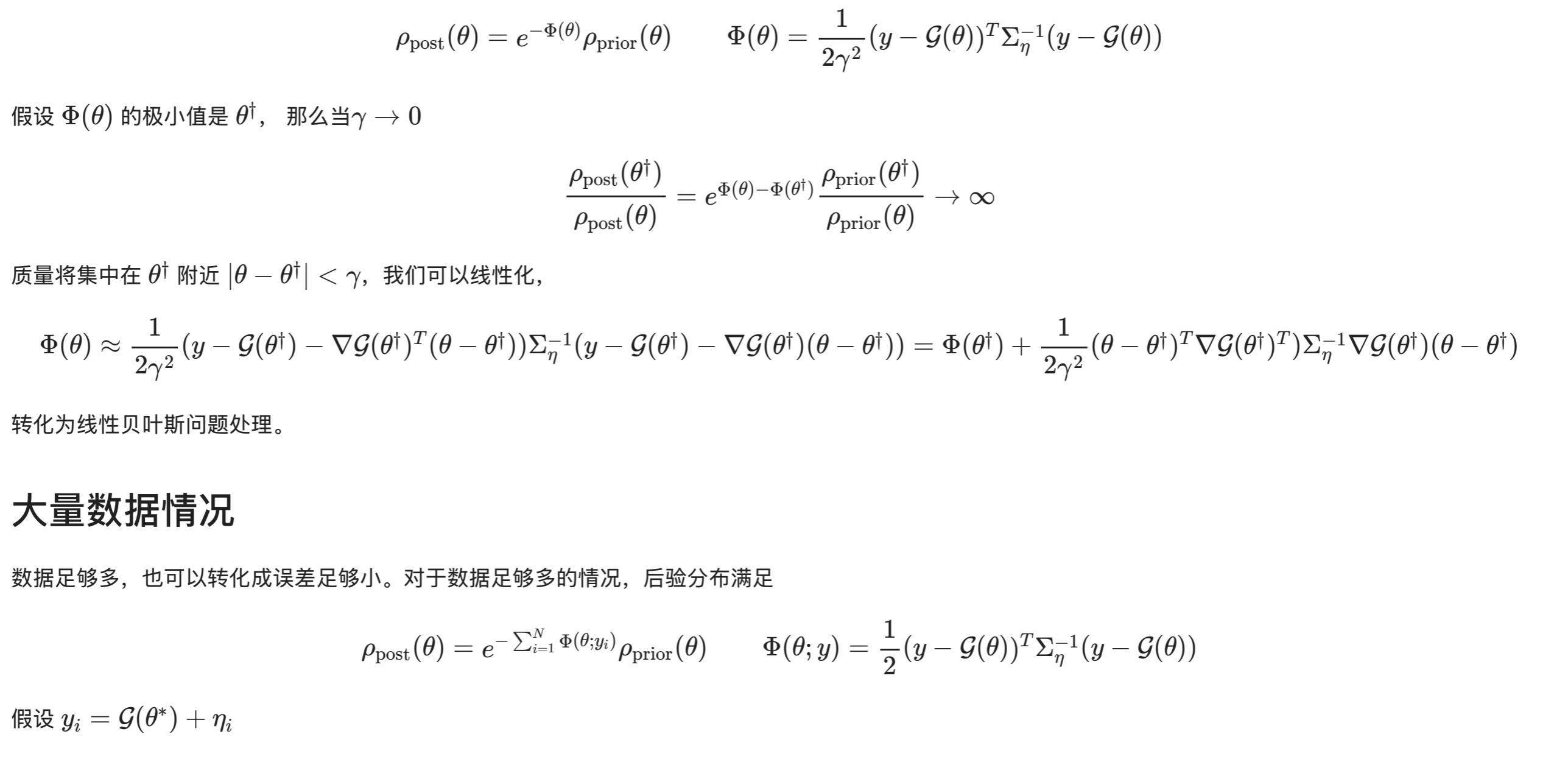
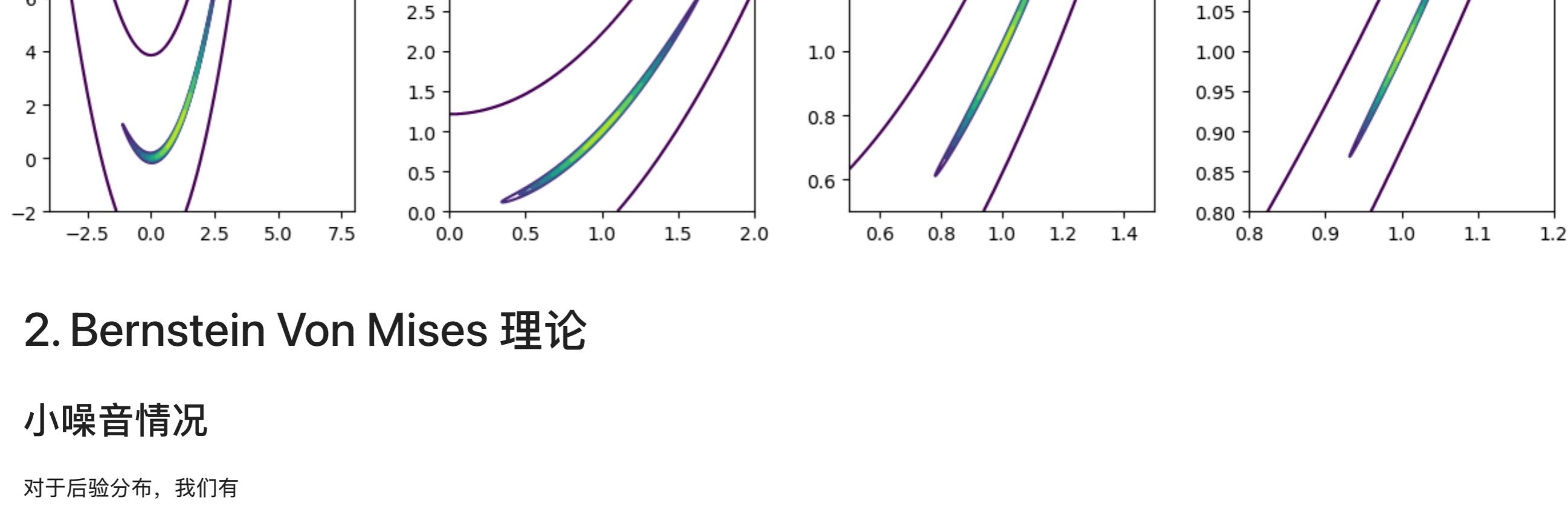
```
using PyPlot

function Phi_R_Multimodal(theta1, theta2, c)
    return ( (4 - (theta2 - theta1)^2)^2 + (c - theta1)^2 + theta2^2 )/2.0
end

Lx, Ux = -3.0, 3.0
Ly, Uy = -3.0, 3.0
N = 1000
X = zeros(N, N)
Y = zeros(N, N)
rho = zeros(N, N)
for i = 1:N
    for j = 1:N
        X[i, j], Y[i, j] = Lx + (Ux - Lx) * (i-1)/(N-1), Ly + (Uy - Ly) * (j-1)/(N-1)
    end
end
rho .= exp.(-rho)
Z = sum(rho)
rho ./= Z
rho ./= (dx * dy)
ax[k].contour(X, Y, rho, 10)
@info "mean = ", sum(X.*rho)/sum(rho), sum(Y.*rho)/sum(rho)
end

fig.tight_layout()
fig.savefig("Multimodal.pdf")
```

```
[ Info: ("mean = ", 0.0, 0.0)
[ Info: ("mean = ", 0.659067998901734, -0.17944598027399047)
```



In []: